Bayesian model selection for Markov chains using sparse probability vectors

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Pink Salmon time series





Year

Pink Salmon time series







Data: Alaska Fisheries Science Center (2018), "AFSC/ABL: Pink salmon data collected at Sashin Creek Weir 1934-2002," NOAA National Centers for Environmental Information, https://inport.nmfs.noaa.gov/ inport/item/17256.

Mixture transition distribution Raftery (1985)

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Example with two states and three lags:

$$Q = \begin{pmatrix} q_{1,1} & q_{1,2} \\ q_{2,1} & q_{2,2} \end{pmatrix} \qquad \begin{array}{c} \text{first order} \\ \text{transition matrix} \end{array}$$

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Sparse Dirichlet mixture (SDM) prior

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$$p(\boldsymbol{\theta} \mid \boldsymbol{n}) \propto \prod_{k=1}^{K} \theta_{k}^{n_{k}} imes \operatorname{Dir}(\boldsymbol{\theta}; \boldsymbol{\alpha}) imes \sum_{k=1}^{K} \theta_{k}^{\beta} \propto \operatorname{Dir}(\boldsymbol{\theta}; \boldsymbol{\alpha} + \boldsymbol{n}) imes \sum_{k=1}^{K} \theta_{k}^{\beta}$$

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Smaller probabilities

Stick-breaking mixture (SBM) prior

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Multinomial data example







 $\Pr(s_t = i_0 \mid s_{t-1} = i_1, s_{t-2} = i_2, \dots, s_{t-R} = i_R, \Lambda, \{\lambda^{(j)}\}_{j=1}^J, \{Q^{(j)}\}_{j=1}^J\} = (\Omega)_{i_R, i_{R-1}, \dots, i_2, i_1, i_0}$

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$$R$$

$$\equiv \Lambda_1 \sum_{\ell=1} \lambda_{\ell}^{(1)} Q^{(1)}(s_t = i_0 \mid s_{t-\ell} = i_\ell) +$$

$$\Pr(s_t = i_0 \mid s_{t-1} = i_1, s_{t-2} = i_2, \dots, s_{t-R} = i_R, \Lambda, \{\lambda^{(j)}\}_{j=1}^J, \{Q^{(j)}\}_{j=1}^J\} = (\Omega)_{i_R, i_{R-1}, \dots, i_2, i_1, i_0}$$

$$\equiv \Lambda_1 \sum_{\ell=1}^R \lambda_\ell^{(1)} Q^{(1)}(s_t = i_0 \mid s_{t-\ell} = i_\ell) + \\ + \Lambda_2 \sum_{\ell_1 < \ell_2} \lambda_{(\ell_1, \ell_2)}^{(2)} Q^{(2)}(s_t = i_0 \mid s_{t-\ell_1} = i_{\ell_1}, s_{t-\ell_2} = i_{\ell_2}) + \ldots +$$

$$\Pr(s_t = i_0 \mid s_{t-1} = i_1, s_{t-2} = i_2, \dots, s_{t-R} = i_R, \Lambda, \{\lambda^{(j)}\}_{j=1}^J, \{Q^{(j)}\}_{j=1}^J\} = (\Omega)_{i_R, i_{R-1}, \dots, i_2, i_1, i_0}$$

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MTD extension

Free parameter count:

K	R	J	Λ	λ	Q	total	unrestricted
5	3	2	1	4	120	125	500
5	4	3	2	11	620	633	$2,\!500$
2	7	4	3	94	30	127	128
3	5	3	2	22	78	102	486
7	5	3	2	22	$2,\!394$	$2,\!418$	$100,\!842$
7	5	2	1	13	336	350	$100,\!842$

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Lag inclusion inference

Dirichlet on mixing parameters



Pink Salmon time series

Lag inclusion inference



Posterior mean of *Q*: Second order with lags 2 and 4 most used

Dirichlet on mixing parameters (.67 for order 2)



Posterior mean of *Q*: Second order with lags 2 and 4 most used

Dirichlet on mixing parameters (.67 for order 2)



SDM on mixing parameters (.86 for order 2)



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MTD extension

Hierarchical representation

$$\begin{aligned} (\boldsymbol{Q}^{(j)})_{i_{j},i_{j-1},...,i_{1},\cdot} &\sim \operatorname{Dir}(\boldsymbol{\alpha}_{\boldsymbol{Q}^{(j)}}) \quad \forall \quad (i_{j},i_{j-1},...,i_{1}) \in \{1,\ldots,K\}^{j}, \\ \boldsymbol{\Lambda} &\sim \operatorname{Dir}(\boldsymbol{\alpha}_{\Lambda}), \qquad \boldsymbol{\lambda}^{(j)} \sim \operatorname{Dir}(\boldsymbol{\alpha}_{\lambda^{(j)}}), \\ \operatorname{Pr}(Z_{t} = j \mid \boldsymbol{\Lambda}) = \boldsymbol{\Lambda}_{j}, \qquad \operatorname{Pr}(\boldsymbol{z}_{t} = (\ell_{1},\ldots,\ell_{j}) \mid Z_{t} = j, \boldsymbol{\lambda}^{(j)}) = \boldsymbol{\lambda}^{(j)}_{(\ell_{1},\ldots,\ell_{j})}, \\ \operatorname{Pr}(s_{t} = i_{0} \mid s_{t-1} = i_{1}, s_{t-2} = i_{2},\ldots,s_{t-R} = i_{R}, Z_{t} = j, \boldsymbol{z}^{(j)}_{t} = (\ell_{1},\ldots,\ell_{j}), \boldsymbol{Q}^{(j)}) = (\boldsymbol{Q}^{(j)})_{i_{\ell_{j}},i_{\ell_{j-1}},\ldots,i_{\ell_{1}},i_{0}} \end{aligned}$$